| Name | Genesis | Notation | p.f. | $\mathbf{E}(\mathrm{X})$ | $\mathbf{V}(\mathbf{X})$ | Applications | Comments |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Uniform (discrete) | Set of $k$ equally likely outcomes (usually, not necessarily, the integers) | $\begin{aligned} & \mathrm{U}(1, \ldots, k) \\ & \text { (not standard) } \end{aligned}$ | $\begin{aligned} & p(x)=1 / k \\ & x=1, \ldots, k \end{aligned}$ | $\frac{k+1}{2}$ | $\frac{k^{2}-1}{12}$ | Dice |  |
| Bernoulli trial | Expt. with two outcomes: 'success' w.p. $\theta$ and 'failure' w.p. $1-\theta$ $X \equiv$ no. successes | $\operatorname{Ber}(\theta)$ | $\begin{aligned} & p(x)=\theta^{x}(1-\theta)^{1-x} \\ & x=0,1 \\ & \theta \in[0,1] \end{aligned}$ | $\theta$ | $\theta(1-\theta)$ | Coins, constituent of more complex distributions |  |
| Binomial | $X \equiv$ no. successes in $n$ ind. $\operatorname{Ber}(\theta)$ trials | $\operatorname{Bi}(n, \theta)$ | $\begin{aligned} & p(x)=\binom{n}{x} \theta^{x}(1-\theta)^{n-x} \\ & x=0,1,2, \ldots, n \\ & \theta \in[0,1] \end{aligned}$ | $n \theta$ | $n \theta(1-\theta)$ | Sampling with replacement | $\operatorname{Bi}(1, \theta) \equiv \operatorname{Ber}(\theta)$ |
| Geometric | $X \equiv$ no. failures until 1st success in sequence of ind. $\operatorname{Ber}(\theta)$ trials | $\mathrm{Ge}(\theta)$ | $\begin{aligned} & p(x)=\theta(1-\theta)^{x} \\ & x=0,1,2, \ldots \\ & \theta \in[0,1] \end{aligned}$ | $\frac{1-\theta}{\theta}$ | $\frac{1-\theta}{\theta^{2}}$ | Waiting times (for single events) | Alternative formulation in terms of $Y \equiv$ no. of trials to 1 st success $(Y=X+1)$ |
| Negative binomial (or Pascal) | $X \equiv$ no. failures to $m$ th success in sequence of ind. $\operatorname{Ber}(\theta)$ trials. Generalization of Geometric | Neg $\operatorname{Bi}(m, \theta)$ <br> (not standard) | $\begin{aligned} & p(x)=\binom{m+x-1}{x} \theta^{m}(1-\theta)^{x} \\ & x=0,1,2, \ldots \\ & \theta \in[0,1] \end{aligned}$ | $\frac{m(1-\theta)}{\theta}$ | $\frac{m(1-\theta)}{\theta^{2}}$ | Waiting times (for compound events) | Neg $\operatorname{Bi}(1, \theta) \equiv \operatorname{Ge}(\theta)$ <br> Remains valid for any $k>0$ <br> (not necessarily integer). <br> Alternative formulation as above. |
| Hypergeometric | $X \equiv$ no. of defectives in sample of size $n$ taken without replacement from population of size $N$ of which $d$ are defective | $\operatorname{Hypergeom}(N, d, n)$ (not standard, esp. order of arguments) | $\begin{aligned} & p(x)=\frac{\binom{d}{x}\binom{N-d}{n-x}}{\binom{N}{n}} \\ & x=\max (0, n+d-N), \ldots \\ & \ldots, \min (n, d) \end{aligned}$ | $\frac{n d}{N}$ | $\frac{N-n}{N-1} n \frac{d}{N}\left(1-\frac{d}{N}\right)$ | Sampling without replacement | Sampling with replacement leads to the $\operatorname{Bi}\left(n, \frac{d}{N}\right)$ - a suitable approx if $\frac{n}{N}<0.1$ |
| Poisson | Arises empirically or via Poisson Process (PP) for counting events. For PP rate $\nu$ the no. of events in time $t \sim \mathrm{Po}(\nu t)$. Also as an approx. to the Binomial | $\operatorname{Po}(\lambda)$ | $\begin{aligned} & p(x)=\frac{e^{-\lambda} \lambda^{x}}{x!} \\ & x=0,1,2, \ldots \\ & \lambda>0 \end{aligned}$ | $\lambda$ | $\lambda$ | Counting events occurring 'at random' in space or time | $\operatorname{Bi}(n, \theta) \equiv \operatorname{Po}(n \theta)$ if $n$ large, $\theta$ small |

SOME CONTINUOUS DISTRIBUTIONS

| Name | Notation | p.d.f. | E(X) | $\mathbf{V}(\mathbf{X})$ | Applications | Comments |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Uniform (continuous) (or Rectangular) | $\mathrm{Un}(\alpha, \beta)$ | $\begin{aligned} & f(x)=\frac{1}{\beta-\alpha} \\ & x \in[\alpha, \beta] \\ & \alpha<\beta \end{aligned}$ | $\frac{\alpha+\beta}{2}$ | $\frac{(\beta-\alpha)^{2}}{12}$ | Rounding errors $U n\left(-\frac{1}{2}, \frac{1}{2}\right)$. <br> Simulating other distributions from $\operatorname{Un}(0,1)$. |  |
| Exponential | $\operatorname{Ex}(\lambda)$ | $\begin{aligned} & f(x)=\lambda e^{-\lambda x} \\ & x>0 \\ & \lambda>0 \end{aligned}$ | $\frac{1}{\lambda}$ | $\frac{1}{\lambda^{2}}$ | Inter-event times for Poisson Process. Models lifetimes of non-ageing items. | Alternative parameterization in terms of $1 / \lambda$ $\operatorname{Ga}(1, \lambda) \equiv \operatorname{Ex}(\lambda)$ |
| Gamma | $\mathrm{Ga}(\alpha, \beta)$ | $\begin{aligned} & f(x)=\frac{\beta^{\alpha} x^{\alpha-1} e^{-\beta x}}{\Gamma(\alpha)} \\ & x \geq 0 \\ & \alpha, \beta>0 \end{aligned}$ | $\frac{\alpha}{\beta}$ | $\frac{\alpha}{\beta^{2}}$ | Times between $k$ events for Poisson Process. Lifetimes of ageing items. | ```Alternative parameterization in terms of \(1 / \beta\) \[ \mathrm{Ga}(1, \lambda) \equiv \operatorname{Ex}(\lambda), \] \[ \mathrm{Ga}(\nu / 2,1 / 2) \equiv X_{\nu}^{2}, \]``` |
| Beta | $\operatorname{Be}(\alpha, \beta)$ | $\begin{aligned} & f(x)=\frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha, \beta)} \\ & x \in[0,1] \\ & \alpha, \beta>0 \end{aligned}$ | $\frac{\alpha}{\alpha+\beta}$ | $\frac{\alpha \beta}{(\alpha+\beta+1)(\alpha+\beta)^{2}}$ | Useful model for variables with finite range. Also as a Bayesian conjugate prior. | $\operatorname{Be}(1,1) \equiv \operatorname{Un}(0,1)$ <br> $\operatorname{Be}(\alpha, \beta)$ is reflection about $\frac{1}{2}$ of $\operatorname{Be}(\beta, \alpha)$. <br> Can transform $\operatorname{Be}(\alpha, \beta)$ on $[0,1]$ to any finite range $[a, b]$ by $Y=(b-a) X+a$ |
| Normal | $\mathrm{N}\left(\mu, \sigma^{2}\right)$ | $\begin{aligned} & f(x)=\frac{1}{\sigma \sqrt{2 \pi}} \exp \left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^{2}\right] \\ & x \in(-\infty, \infty) \end{aligned}$ | $\mu$ | $\sigma^{2}$ | Empirically and theoretically (via CLT etc.) a good model in many situations. Often easy to handle mathematically. | $\begin{aligned} & X \sim \mathrm{~N}\left(\mu, \sigma^{2}\right) \Longrightarrow \\ & a X+b \sim \mathrm{~N}\left(a \mu+b, a^{2} \sigma^{2}\right) \\ & \Longrightarrow Z=\frac{X-\mu}{\sigma} \sim \mathrm{N}(0,1) \\ & \mathrm{So} \end{aligned}$ $P[X \in(u, v)]=P\left[Z \in\left(\frac{u-\mu}{\sigma}, \frac{v-\mu}{\sigma}\right)\right]$ <br> $\mathrm{N}(0,1)$ special case has p.d.f. denoted $\phi$, c.d.f. $\Phi$ (tabulated). <br> Note $\Phi(-z)=1-\Phi(z)$. |
| Chi-square | $\chi_{\nu}^{2}$ | $\begin{aligned} & f(x)=2^{-\nu / 2} \Gamma(\alpha)^{-1} x^{\nu / 2-1} e^{-x / 2} \\ & x>0 \\ & \nu>0 \end{aligned}$ | $\nu$ | $2 \nu$ | Sum of squares of $\nu$ standard normals | $X_{\nu}^{2} \equiv \mathrm{Ga}(\nu / 2,1 / 2)$ $\text { If } X_{1}, X_{2}, \ldots, X_{n} \sim \mathrm{~N}(0,1)$ <br> independent, then $\sum_{i=1}^{n} X_{i}^{2} \sim \chi_{n}^{2}$ |
| Student $t$ | $\mathrm{t}_{\nu}$ | $\begin{aligned} & f(x)= \\ & \nu^{-1 / 2} B\left(\frac{1}{2}, \frac{\nu}{2}\right)^{-1}\left(1+x^{2} / \nu\right)^{-(\nu+1) / 2} \\ & x \in(-\infty, \infty) \\ & \nu>0 \end{aligned}$ | $\begin{aligned} & 0 \\ & \text { (if } \stackrel{1}{\nu} \text { ) } \end{aligned}$ | $\frac{\nu}{(\mathrm{if} \stackrel{\nu}{\nu}}$ | Useful alternative to Normal for variables with heavy tails. | If $X \sim \mathrm{~N}(0,1)$ and $Y \sim \chi_{\nu}^{2}$ independent then $\begin{aligned} & \frac{X}{\sqrt{Y / \nu}} \sim \mathrm{t}_{\nu} . \\ & \mathrm{t}_{1} \equiv \text { Cauchy. } \mathrm{t}_{\nu}^{2} \equiv \mathrm{~F}_{1, \nu} . \end{aligned}$ |
| F | $\mathrm{F}_{\nu, \delta}$ | $\begin{aligned} & f(x)=\frac{\nu^{\nu / 2} \delta^{\delta / 2} x^{\nu / 2-1}}{B(\nu / 2, \delta / 2)(\nu x+\delta)^{(\nu+\delta) / 2}} \\ & x>0 \\ & \nu, \delta>0 \end{aligned}$ | $\begin{gathered} \frac{\delta}{\delta-2} \\ (\text { if } \delta>2 \text { ) } \end{gathered}$ | $\begin{aligned} & \frac{2 \delta^{2}(\nu+\delta-2)}{\nu(\delta-2)^{2}(\delta-4)} \\ & (\text { if } \delta>4) \end{aligned}$ | Scaled ratio of chi-squares. Used in tests to compare variances | If $X \sim \chi_{\nu}^{2}$ and $Y \sim \chi_{\delta}^{2}$ independent then $\frac{X / \nu}{Y / \delta} \sim \mathrm{F}_{\nu, \delta}$. <br> If $T \sim \mathrm{t}_{\nu}$ then $T^{2} \sim \mathrm{~F}_{1, \nu}$. <br> If $Z \sim \operatorname{Be}(\alpha, \beta)$ then $\frac{\beta Z}{\alpha(1-Z)} \sim \mathrm{F}_{2 \alpha, 2 \beta}$. |

